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# Understanding the STEM Pipeline

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WORKING PAPER 125 · JANUARY 2015

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# **Acknowledgements**

I wish to thank the staff of the Florida Department of Education's Education Data Warehouse for their help in obtaining and interpreting the data used in this study. Funding for this research was provided by the Gates and Smith Richardson Foundations as part of a grant to the Center for the Analysis of Longitudinal Data in Education Research (CALDER) and American Institutes for Research. Any opinions or errors are solely attributable to me, however.

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The views expressed are those of the authors and should not be attributed to the American Institutes for Research, its trustees, or any of the funders or supporting organizations mentioned herein. Any errors are attributable to the authors.

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Tim R. Sass CALDER Working Paper No. 125 January 2015

# Abstract

I investigate the determinants of high school completion and college attendance, the likelihood of taking science, technology, engineering or math (STEM) courses in the first year of college and the probability of earning a degree in a STEM field. The focus is on women and minorities, who tend to be underrepresented in STEM fields. Tracking four cohorts of students throughout Florida, I find that large differences in math achievement across racial lines exist as early as elementary school and persist through high school. These achievement differences lead to higher drop-out rates in high school and a reduced probability of attending college for black students. However, conditional on immediately attending a four-year college after high school, black and Hispanic students are more likely than whites to take STEM courses during their first year in college. Increased exposure to Hispanic math and science teachers in middle and high school tends to increase the likelihood that Hispanic students take STEM courses during their first year in college, though pairing black students and black math/science teachers does not have the same positive effect. For all students, having high school math and science teachers with a degree in biology, chemistry or math (as opposed to education) is associated with a higher likelihood of taking STEM courses as college freshmen. When pre-college differences in income and math achievement are taken into account, black and Hispanic students are at least as likely as white students to successfully complete a STEM major. Racial/ethnic pairing of students and college instructors in first-year STEM courses does not increase the likelihood of majoring in a STEM field. In contrast to underrepresented minorities, women perform nearly as well as men on math achievement tests through high school and are more likely to finish high school and attend college than males. Among college students, however, women are less likely than men to take courses in the physical sciences in their first year and are less likely to earn a degree in physics or engineering, even after adjusting for pre-college test scores. Gender matching of students and math/science teachers in middle and high school tends to increase the likelihood that female college freshman will take at least one STEM course, However, conditional on first-year coursework, neither gender matching at the secondary or college levels appears to have any effect on the likelihood of completing a major in a STEM field.

## I. Introduction

There is growing concern that the United States does not produce a sufficient number of majors in science, technology, engineering and math (STEM) fields in order to remain globally competitive. Of particular concern is the underrepresentation of women and minorities in STEM fields. Various hypotheses have been forwarded to explain why there are relatively few women and minorities in STEM fields. These include negative peer effects of male students in math and science courses, lack of same gender/same race instructors as role models and the extent of institutional focus on undergraduate education.

Most of the existing evidence on enrollment and persistence in STEM majors is based on experiences while in college. However, this may obscure important influences of student preparation and experiences in high school and early grades. The focus on courses, instructors and grades while in college is primarily a matter of data availability. Most extant studies have relied on college transcript data which provide little information about what has happened to a student prior to attending college.

The present study of STEM degree attainment is the first to employ K-12 data linked to college transcripts in order to track individual students all the way from elementary school through the end of college. These data allow one to determine where gaps emerge between males and females and between whites and under-represented minorities. Such information is important in determining where to target interventions designed to promote female and minority participation in STEM fields. Indeed, the evidence reveals that gaps in math achievement between blacks and whites occur as early as elementary school and have substantial effects on high school completion, college attendance and ultimately STEM major completion. In contrast, women perform nearly as well as men on math achievement tests through high school, are more likely to finish high school and attend college than males. Among college students, however, women are much less likely than men to earn a degree in a STEM field, even after adjusting for prior test scores. There is some evidence that gender and ethnic matching of students and secondary teachers increases the likelihood that women and Hispanics take STEM classes as college freshmen. Likewise, taking high school high school courses from teachers who possess a baccalaureate degree in math or science (rather than education) is associated with a higher probability of taking STEM courses as a college freshman. However, conditional on first-year college coursework, gender or ethnic/racial matching of students and college instructors is not associated with completing a degree in a STEM field.

## II. Existing Evidence

Spurred by the intense policy interest in the underrepresentation of women and minorities in STEM fields, there has been a recent surge in research on the determinants of entry and persistence in STEM majors. Most of this recent work focuses on major choice conditional upon enrolling in college.

#### A. STEM Persistence Among College Students

Bettinger and Long (2005) employ individual level data on students who entered four-year public colleges and universities in Ohio in 1998 and 1999. They find that exposure to a female college instructor in their first college course in a subject has mixed effects on female students' participation in STEM. Female students who initially had a female instructor were less likely to take additional courses in biology and physics than similar students whose first professor in the subject was male. However, initial exposure to a female instructor boosted the likelihood that a female student would take an additional course in the subject and the total number of hours of credit hours in geology, mathematics and statistics. They found no effects of instructor gender on the major choices of female college students in STEM fields. In later work, Bettinger (2010), finds STEM fields only retain about one-half of students. Further, the proportion of first-year courses that are taken in STEM fields is directly correlated with an eventual major in STEM. Price (2010) analyzes the same Ohio data, but for a longer time period, 1998-2002. He focuses on the impact of race and gender matching of students and instructors in first-semester freshman STEM courses on student persistence in STEM majors. He finds that black students are more likely to persist in a STEM major if they have a STEM course taught by a black instructor. As in Bettinger and Long, female students are less likely to persist the greater the proportion of their STEM courses which are taught by female instructors.

Hoffman and Oreopoulos (2009) use data from the University of Toronto to analyze the impact of instructor gender in large first-year undergraduate courses on persistence within a course and later course taking in the same subject. With multiple subjects they are able to employ student fixed effects to control for unobserved student heterogeneity. In the fixed-effects models they find that gender matching of students and instructors has no effect on the likelihood of dropping the course, but does increase the number of additional courses taken in the subject. This later effect does not hold for math and science courses, however, where gender matching actually reduces the number of additional courses taken in the subject.

Fairlie, Hoffman and Oreopoulos (2014) also conduct a single-institution analysis, but focus on racial matching of students and instructors and employ data from a California community college, rather than a four-year university. They find that under-represented minority students are less likely to drop a course, more likely to earn a "B" or better if they complete the course and more likely to take an additional course in the same subject the following term when taught by an under-represented minority instructor. These effects tend to be stronger when there is an exact match between the race/ethnicity of the student and instructor. The authors also uncover evidence that racial matching increases the likelihood a student will remain in college for two quarters and boosts the probability an under-represented minority student obtains an associates or vocational degree. The impact of student-instructor racial matching on the likelihood of transferring to a four-year institution is mixed.

Kokkelenberg and Sinha (2010) analyze graduation probabilities of students at SUNY-Binghamton. They find that student SAT math scores and having AP credits in a STEM field are positively correlated with completing a major in a STEM field. However, conditional on AP credits and SAT scores, women, blacks and Hispanics are all less likely than whites to graduate with a major in a STEM field.

Griffith (2010) finds that students at selective research institutions are less likely to persist in STEM fields than their counterparts in lower ranked colleges and universities. Like Bettinger and Long, Griffith finds a zero or negative correlation between percent of female faculty in STEM and female persistence in a STEM major. However, for both women and minorities, the proportion of STEM PhDs who are in the same gender/racial category is positively correlated with STEM persistence among undergraduates.

Consistent with Griffith, Arcidiacono, et al. (2013) finds that students sort into majors based on academic preparation (high school GPA, SAT scores). Among schools in the University of California (UC) system, lower ranked schools were better at graduating less-prepared minority students. Most minorities were more likely to graduate with a science degree if they attended a lower-ranked UC school.

Carrell, Page and West (2010) analyze the effects of instructor gender on student achievement, course taking and STEM major choice at the U.S. Air Force Academy. The Air Force Academy requires all students to take a set of standardized introductory courses and randomly assigns students to classrooms, thereby avoiding any possible selection bias associated with student course and instructor choice. In contrast to other studies, they find substantial positive effects of female instructors on female student's performance in math and science courses at the Academy. Further, high-performing young women, as measured by SAT math scores, are more likely to take additional STEM courses and graduate with a degree in a STEM major if they have introductory courses from a female instructor. For these high-ability women who take introductory courses from a female instructor, the STEM gender gap is

eliminated. It is unclear, however, whether their strong gender matching findings are due to their rigorous experimental design or are related to unique aspects of the female students at an elite military academy.

#### B. Pre-College Influences on Educational Attainment in STEM Fields

The literature relating pre-college experiences to college outcomes is much thinner. In a wellknown set of papers, Dee finds student teacher matching has positive effects on pre-college outcomes. Dee (2004), using data from the STAR experiment in Tennessee finds that assignment to an own-race teacher had positive effects on student achievement in math in early elementary grades. Dee (2007), employing a cross-subject student fixed effects strategy, finds that 8<sup>th</sup> grade female students exposed to a female science teacher had insignificant effects on science test scores and assignment to a female math teacher significantly lowered math achievement for girls. However, Dee's evidence ends at the middle school level and provides no direct evidence on outcomes in high school and college.

Park, Behrman and Choi (2012) exploit the random assignment of students to high schools in Korea to estimate the causal effects of same-sex schools on math college entrance exam scores, student expectations of college attendance and major choice while in high school and actual college attendance and major choice two years after high school. While attending an all-girls school has positive effects on entrance exam scores in general math and calculus, no effects were found on interest in science, student expectations or actual choice of a STEM major.

Anelli and Peri (2013) link data on graduates of college-preparatory high schools in Milan to enrollment records of universities in Milan in order to study the effects of high school gender composition on college major choice. In the Italian system high school students are grouped into "classes" during their freshman year and maintain these groupings throughout high school. Rather than study STEM fields, they divide majors into "high earning" and "not high earning" categories and consider the determinants of initial major choice. They find that an increase in the proportion of female

classmates in high school classes is associated with an increased likelihood of women selecting a high earning major.

## III. Data

The data for this study come from a variety of sources. The primary source for student-level information is the Florida Department of Education's K-20 Education Data Warehouse (K-20 EDW), an integrated longitudinal data base covering all public school students and teachers in the state of Florida. For K-12 students, the K-20 EDW provides demographic information, enrollment and attendance, program participation, disciplinary actions and achievement test scores, beginning in 1995. Florida began testing students statewide in 1997/98, with the introduction of the "Sunshine State Standards" Florida Comprehensive Achievement Test (FCAT-SSS). The FCAT-SSS is a criterion-based exam designed to test for the skills that students are expected to master at each grade level. It is a "high-stakes" test used to determine schoolgrades, student retention in some grades and passage of the 10<sup>th</sup> grade exam is a requirement for graduation from high school. The FCAT-SSS was initially administered to selected grades but was later expanded to grades 3-10 in 2000/01. Beginning in 1999/2000, a second test, the FCAT Norm-Referenced Test (FCAT-NRT) was added in each of grades 3-10. The FCAT-NRT, was a custom form of the Stanford Achievement Test used throughout the country. No accountability measures were tied to student performance on the NRT. Florida stopped administering the FCAT-NRT after 2007/08. The FCAT-SSS exam was replaced with the FCAT 2.0 beginning in 2010/11.

As the name implies, the K-20 EDW also includes records for students enrolled in community colleges or four-year public universities in Florida. The K-20 EDW also contains information on the Florida Resident Assistance Grant (FRAG), a grant available to Florida residents who attend private colleges and universities in the state. Data from the National Student Clearinghouse (NSC), a national database that includes enrollment data from 3,300 colleges throughout the United States, is used to

track college attendance outside the state of Florida, as well as any private college enrollment in Florida that the FRAG data do not pick up. Unfortunately, the Florida Department of Education's data-sharing agreement with the NSC expired in the latter part of the 2000s, so I can only reliably track students who attended private colleges and universities within Florida or any postsecondary institution outside of Florida through school year 2006–2007.<sup>1</sup> Enrollment, coursework and degree attainment information are available for all post-secondary students at public institutions in Florida. In addition, demographic information on post-secondary instructors is available as well.

High school graduation is determined by withdrawal information and student award data from the K-20 EDW. While various diploma options exist, including a GED and a special-education diploma, I focus on receipt of a regular high school diploma. Students who withdrew with no intention of returning or exited for other reasons such as non-attendance, court action, joining the military, marriage, pregnancy, and medical problems, but did not later graduate, are counted as dropouts. It is not possible to directly determine the graduation status of students who leave the Florida public school system to attend a home-schooling program, to enroll in a private school or who move out of state.

The analysis sample covers four cohorts of 5<sup>th</sup> grade students. Statewide achievement testing for 5<sup>th</sup> grade students began in the 1997/98 school year, so the first cohort in the sample are students who attended 5<sup>th</sup> grade for the first time in 1997/98 and took the FCAT-SSS math exam. The final cohort is composed of students who were enrolled in 5<sup>th</sup> grade for the first time in 2000/01. Descriptive statistics for these four cohorts of students are provided in Table 1.

The last year for which student data are currently available is 2012/13. Given that it takes at least three years to progress through middle school and high school completion typically takes four years, this means the each of the four cohorts can be tracked through high school and into the

<sup>&</sup>lt;sup>1</sup> Information on the NSC is available at www.studentclearinghouse.org.

beginning of college. If we allow five years for college completion, all four cohorts can be followed all the way through the end of college. In order to analyze baccalaureate completion, it is therefore necessary to restrict the sample to students who do not repeat any grades in middle school. Descriptive statistics for this restricted sample are provided in Table 2. As one would expect, students who follow the normal progression through middle school have much higher test scores than those who repeat a grade or drop out before completing high school and thus the test score means in Table 2 are higher than those in Table 1.

# IV. Analysis and Results

#### A. Descriptive Analysis of Pre-College Outcomes and College Attendance

In order to understand where the "leakages" in the STEM pipeline occur, I begin with a descriptive analysis of achievement differences in math prior to high school entry. Tables 3A, 3B and 3C report mean test scores in elementary and middle school by race, gender and school lunch status, respectively. The differences between blacks and whites are striking. White students outperform black students in math by roughly two-thirds of a standard deviation and this differential is essentially constant across grades. The differences between Hispanics and Anglos are less pronounced. Hispanic students tend to score around the mean, whereas Anglos score 0.3 to 0.4 standard deviations above the mean. As with blacks, the differences between Hispanics and Anglos are relatively constant throughout elementary and middle school.

Achievement scores across gender tell a much different story. Whereas large and persistent gaps between racial/ethnic groups persist throughout elementary and middle school, any differences in math achievement for boys and girls are relatively small and inconsistent. Based on the FCAT-NRT (i.e. the Stanford Achievement Test) it appears that girls start out roughly on par with boys in 5<sup>th</sup> grade and continue there through 8<sup>th</sup> grade. Results are less stable for the FCAT-SSS exam; girls start out slightly

below boys in 5<sup>th</sup> grade (about 0.04 standard deviations) and end up slightly above in 8<sup>th</sup> grade (about 0.04 standard deviations). The average test scores broken down by school lunch status (a measure of poverty) mirror those by race. Children who receive lunch for free score about 0.7 standard deviations below their counterparts who receive no lunch subsidy and the achievement gap remains constant over time. Achievement levels of children who receive reduced-price lunches (and thus have family incomes above those of free-lunch recipients) are in-between those of free-lunch recipients and students who do not qualify for any lunch subsidy.

The high school experiences of students, broken down by race, gender and family income, play out much like the achievement measures in elementary and middle school. Tables 4A, 4B and 4C provide information on exit propensities by race/ethnicity, gender and family income, respectively. Blacks and Hispanics are much more likely to drop out of high school (approximately 10 percent), than whites or Asians (6 and 3 percent, respectively). These drop-out rates are conditional on students progressing through middle school in three years. The disparities when students who repeat one or more middle school grades are included (not reported) are even higher. As with test scores prior to high school, the differences in drop-out rates between boys and girls are relatively small, with females about 1 percentage point less likely to drop out of high school. High school drop-out rates across income categories mirror those by race; about 12 and one-half percent of free-lunch students drop out of high school while only 4 percent of students with no lunch subsidy drop out.

Even if students complete five years of high school, they may not earn a regular high school diploma. They could receive a GED or (if they are a special education student) a special diploma or certificate of completion. Alternatively, they could remain enrolled, but still not have obtained a diploma within five years. As demonstrated in Tables 5A and 5C, there are large differences in receipt of a regular high school diploma between blacks and whites and between students from poor families and those from more affluent families. About 58 percent of black students earn a regular high school

diploma within 5 years of entering high school whereas over 70 percent of white students obtain a regular high school diploma. Asian students do better than any other racial/ethnic group; over 85 percent of Asian students obtain a regular high school diploma with five years. As with test scores and drop-out rates, diploma attainment differences across family income levels mimic differences between blacks and whites. If there is a gender gap in obtaining a regular high school diploma, it is in favor of females. As shown in Table 5B, the diploma-receipt rate for young men is 64 percent compared to 70 percent, for young women.

A final measure of high school performance is achievement test scores. Tables 6A-6C provide average math test scores broken down by race/ethnicity, gender and family income, respectively. Large achievement gaps between blacks, whites and Hispanics remain essentially constant all the way from 5<sup>th</sup> through 10<sup>th</sup> grade, even though students who drop out in the first two years of high school are removed from the sample. Similarly, significant achievement gaps across family income categories (proxied by lunch status) remain relatively constant from elementary through high school. The achievement gap between boys and girls also persists through high school, though is much smaller than disparities across race or income. Using the FCAT-NRT scores as a metric, boys outperform girls in 5<sup>th</sup> grade by 0.025 standard deviations and this widens slightly to 0.041 standard deviations in 10<sup>th</sup> grade. In contrast, scores from the FCAT-SSS suggest a slight narrowing in the math achievement gap between boys and girls as they progress through school; the achievement differential is 0.062 standard deviations in 5<sup>th</sup> grade and only 0.044 by grade 10. While the gaps are small in either case, the differential findings across tests may reflect the fact that the NRT exam tends to test higher-level math skills than the SSS exam.

Conditional on earning a regular high school diploma, we observe substantial differences in post-secondary educational choices by race/ethnicity, gender and income. As shown in Table 7A, Blacks are much less likely than whites to attend any college whereas the college-going rate for Hispanics is

comparable to that of whites.<sup>2</sup> Similarly, a reported in Table 7C, the likelihood of attending either a twoyear or four-year college increases with family income. Table 7B documents the fact that the college going rates of women are much higher than men (72 percent versus 64 percent).

Taken together, the descriptive evidence shows that for minorities and students from lowincome families, there are substantial "leakages" in the educational pipeline before college. For women, however, their educational attainment and test scores would suggest they would be as likely as men to complete a STEM major in college. I turn now to analyses of course selection, persistence and major choice among those students who make it to college.

#### B. Determinants of Initial College Coursework

Table 8 presents probit estimates of the probability of taking one or more STEM courses in the first year of college, conditional upon having earned a regular high school diploma within five years of starting grade 9 and enrolling in a four-year public university within a year of receiving their high school diploma. The first column reports estimates with only student race/ethnicity in the model. Without any other controls, black and Hispanic students are more likely than whites (the omitted group) to take at least one STEM course during their first year in college.<sup>3</sup> In contrast, women are about 4 percentage points less likely than men to take at least one STEM course in their freshman year. Controlling for family income (5<sup>th</sup>-grade lunch status) and math test scores in middle and high school does not alter these results.

As shown in Table 9, the relationships between STEM course taking and race/ethnicity and gender remain largely the same when considering specific types of science courses. Conditional on pre-

 $<sup>^2</sup>$  While it appears that blacks are more likely to attend out-of-state colleges, this is likely due to a peculiarity of the National Student Clearinghouse Data. Universities with nationwide branch campuses, like the University of Phoenix are all assigned to the state of their headquarters, even if they have a physical presence in another state.

<sup>&</sup>lt;sup>3</sup> The level of courses is not taken into account, so some math and sciences could be remedial in nature.

college test scores and family income, both blacks and Hispanics who start college at a four-year public university are more likely to take at least one science course during their freshman year and more likely to take a biology or chemistry course as a freshman than are white (non–Hispanic) students. Women are less likely than men to take any science course and less likely to take physics during their freshman year. They are more likely than men, however, to take at least one biology course during their first year of college.

In Table 10 I present estimates of the relationship of the characteristics of middle and high school teachers to first-year college course selection.<sup>4</sup> For Hispanic students who attend a four-year public university immediately after completing high school, increases in the proportion of Hispanic secondary math and science teachers they were exposed to are associated with a greater likelihood they will take at least one STEM course as college freshmen. Likewise, increases in the proportion of middle and high school math courses that were taught by female teachers is positively correlated with the likelihood a female student will take at least one STEM course during their first year in college. These positive matching effects do not hold for black students, however. Increases in exposure to black middle and high school science and math teachers are in fact negatively correlated with the probability of taking at least one STEM course as a college freshman.

While there is some evidence that gender or ethnic matching of students and teachers in secondary school impacts initial college course selection, the preparation of math and science teachers matters as well. Having a biology, chemistry or math course in high school taught by a teacher with a bachelor's degree in the relevant subject (rather than a math or science education degree or a degree in

<sup>&</sup>lt;sup>4</sup> Note that I control for pre-high school math test scores, but not scores in grades 9 and 10, so the high school teacher effects could be working through effects on high school math achievement.

another science field) has small but statistically significant effects on the likelihood a student later takes at least one STEM course as a freshman in college.<sup>5</sup>

As illustrated in Table 11, the link between preparation of high school teachers and specific STEM course selection for college freshman is less clear. Having had a high school chemistry course taught by a teacher with a BA degree in chemistry is associated with an increased likelihood of taking chemistry as a college freshman, but the same is not true for biology or physics. Having at least one high school math course taught by at teacher with a BA in math appears to boost the likelihood of taking biology, chemistry or at least one science course during the first year in college.<sup>6</sup>

### C. Determinants of Completing a College Degree in a STEM Major

Table 12 presents results of estimating probit equations which predict completion of college majors (conditional on attending a Florida public university immediately after earning a high school degree), without any controls for pre-college family income or student achievement test scores. These estimates reveal the expected pattern; black and Hispanic college students are less likely to earn a degree in a STEM field than are white (non-Hispanic) students. Likewise, women are less likely than men to successfully complete a major in a STEM field. These patterns generally hold across specific STEM majors. The one notable exception is Biology and other life sciences, where women have a higher likelihood of earning a bachelor's degree than men.

Table 13 presents estimates of STEM major completion in which family income (proxied by free/reduced-price lunch status in grade 5) and middle and high school math achievement scores are included as controls. Holding constant family income and prior test scores, black college students are no

<sup>&</sup>lt;sup>5</sup> I find no effects for physics teachers with a BA degree in physics, but this may simply be due to the very small number of high school physics teachers in Florida who possess a degree in physics.

<sup>&</sup>lt;sup>6</sup> One possibility is that the college majors of high school math and science teachers are simply reflecting the general quality of a high school and not direct effects of teacher preparation on student interest in math and science.

less likely than whites to major in a STEM field and Hispanics are slightly more likely to successfully complete a major in a STEM field than Anglos. For a variety of specific STEM majors, there are no differences in degree attainment between white, black and Hispanic students conditional on family income and prior math achievement. Put differently, if students make it to college and have equal resources and pre-college preparation, I find no differences across race/ethnicity in the likelihood of earning a bachelor's degree in a STEM field.

In general, conditional on college entry, family income and prior math achievement, there is no significant difference between women and men in the probability of graduating with a degree in a STEM field. However, there are significant gender differences in the likelihood of degree attainment across STEM majors. Women are more likely than men to complete a major in life sciences, but are less likely to earn a degree in physics or engineering than men.

To explore possible differences in the likelihood of persisting in a major, I estimate models of degree completion conditional on first-year course selection. Results are presented in Tables 14 and 15. Not surprisingly, freshman-year coursework is strongly related to eventual degree attainment. The greater the number of freshman-year engineering courses taken the more likely a student will eventually earn an engineering degree. The same is true for math courses and math degrees and freshman science courses and degrees in chemistry, physics and life sciences. Holding constant first-year coursework, black students and female students are less likely to complete a degree in a STEM field than are white males. In general, the probability of an Hispanic student persisting in a STEM field (conditional on family income and pre-college math test scores) is no different than that of an Anglo student.

In order to determine if student-teacher matching by race/ethnicity or gender effects persistence in a STEM major I add controls for both middle/high school and first-year-in-college matching of students and instructors. Estimates of this expanded model of degree attainment

conditional on initial college course work are presented in Table 15. While student-teacher matching in middle and high school math and science courses appeared to influence freshman-year course selection, boosting freshman-year STEM course taking for women and Hispanics (see Table 10), the same is not true for completion of a STEM major conditional on first-year course work. In fact the correlation between racial/ethnic and gender matching of students and teachers in middle and high school and STEM major persistence is frequently negative.

# V. Summary, Conclusions and Future Research

Growing concern about the low production of college graduates in STEM fields, particularly among minorities and women, has led to a rapidly growing research literature seeking to understand the causes of these disparities and hence provide guidance as to appropriate policies. The focus of this research has been on the experiences of students once they attend college, including the identity of their instructors and the institutions they enroll in. This college-level focus ignores the fact that students have to get to college first before they can even stand a chance at obtaining a college degree in a STEM field. It also forecloses the possibility that pre-college experiences, such as the quality and identity of middle and high school instructors can shape future major choices in college.

In this paper, I present new data tracking individual students from elementary school through college. I find that for blacks and students from low-income families, achievement gaps begin early and persist through high school. These gaps are associated with higher drop out, reduction in the probability of obtaining a regular high school diploma and decreased likelihood of attending college. However, black and Hispanic students who enter a four-year college immediately after high school are more likely than white students to take at least one STEM course during their freshman year. Increases in the proportion of Hispanic math and science teachers in middle and high school are positively correlated with the number of first-year STEM courses taken by Hispanic college freshmen. Pairing of black

students with black middle and high school math and science teachers does not boost the number of first-year STEM courses taken by black college students, however. For all students, exposure to biology, chemistry and math high school teachers who possess a bachelor's degree in the relevant subject area is associated with increased STEM course taking as college freshmen. While black and Hispanic college students are less likely to complete a major in a STEM field than their white counterparts, if income and pre-college achievement differences are taken into account, both blacks and Hispanics who make it to college are no less likely than whites to graduate with a degree in a STEM field. Conditional on first-year course selection, neither the identity of secondary math and science teachers nor the identity of college instructors in first-year college courses appears to alter the likelihood that underrepresented minorities will obtain a bachelor's degree in a STEM field.

Gender gaps in math achievement are generally small throughout elementary, middle and high school and women are more likely to successfully complete high school and attend college. Once they get to college, however, they are much less likely than males to obtain a bachelor's degree in a STEM field. Although female college students are more likely to complete a major in biology or other lifescience fields, they are much less likely than men to earn a degree in engineering or the physical sciences. Exposure to female math and science teachers in middle and high school are correlated with increases in the number of STEM courses taken by female college freshmen. However, the gender matching of students and teachers in secondary school does not increase persistence of women in STEM fields after their freshman year. Similarly, conditional on first-year coursework, exposure to female instructors in STEM courses taken during the freshman year does not increase the likelihood of successfully completing a major in a STEM field.

The analyses conducted thus suggest that altering the gender and racial composition of college faculty are unlikely to substantially change the relative numbers of women and underrepresented minorities in STEM fields. However, increasing the proportion of female math and science teachers in

high school and hiring more high school math and science teachers with degrees in the relevant subject area are likely to boost the number of STEM courses taken by both women and under-represented minorities during their first year in college. The increased first-year coursework in turn will tend to boost the numbers of women, blacks and Hispancis who successfully complete a STEM major. Further, policies aimed at reducing racial achievement gaps in the early grades are also likely to enhance the representation of blacks in STEM fields in the long run.

One potential drawback of the present analysis is that the identities of high school and college instructors have been treated as exogenous. At both the secondary and college levels students have some discretion over the identity of the instructors they take courses from and institutions make choices about the composition of their faculties. Thus student-teacher matching effects could be biased by student self-selection or unmeasured attributes of educational institutions. To alleviate this potential problem I will pursue two strategies in future work. One approach will be to include high school fixed effects, which should account for unmeasured factors that may be correlated with the gender and racial/ethnic composition of high school faculties. The second approach, which has been employed in previous work by others, will be to use the overall faculty composition as an instrument for the matching of individual students and teachers.

The next phase of the research will be to link data on classroom composition at both the K-12 and post-secondary levels to individual students in math and science courses. This will allow tests of peer effects on both college coursework and major choices. I can also expand the sample to include students who start their college careers at community colleges to see if the patterns found in universities also hold in two-year institutions.

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# Tables

# Table 1 – Descriptive Statistics by 5<sup>th</sup>-Grade Cohort

(Students with a test score in grade 5 who are enrolled three or more years later in a public school in Grade 9)

		Coho	rt	
	1997	1998	1999	2000
Female	0.5000	0.4971	0.4954	0.4982
White	0.5292	0.5366	0.5211	0.5119
Black	0.2469	0.2490	0.2483	0.2409
Hispanic	0.1972	0.1847	0.1996	0.2138
Asian	0.0196	0.0215	0.0214	0.0216
Race-other	0.0072	0.0082	0.0097	0.0118
Free Lunch	0.4253	0.4125	0.4146	0.4082
Reduced-Price Lunch	0.0931	0.1036	0.1048	0.1093
5th-grade SSS Normed Math Score	0.0050	0.0075	0.0151	0.0199
5th-grade NRT Normed Math Score			0.0156	0.022
6th-grade SSS Normed Math Score	-0.5011	-0.5446	0.0661	0.0809
6th-grade NRT Normed Math Score	-0.5981	0.0721	0.0638	0.0779
7th-grade SSS Normed Math Score	-0.5654	0.0850	0.0735	0.082
7th-grade NRT Normed Math Score	0.0590	0.0736	0.0695	0.079
8th-grade SSS Normed Math Score	0.0636	0.0801	0.0688	0.0782
8th-grade NRT Normed Math Score	0.0515	0.0703	0.0564	0.067
9th-grade SSS Normed Math Score	0.1327	0.1269	0.1233	0.115
9th-grade NRT Normed Math Score	0.1199	0.1081	0.1045	0.096
10th-grade SSS Normed Math Score	0.0914	0.2122	0.1828	0.200
10th-grade NRT Normed Math Score	0.0566	0.0637	0.0551	0.073
Earned Regular HS Diploma within 4 years of Entering Grade 9	0.6299	0.6095	0.6213	0.642
Earned Regular HS Diploma within 5 years of Entering Grade 9	0.6323	0.6121	0.6247	0.646

Table 2 - Descriptive	Statistics by	5th-Grade	Cohort
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(Students with a test score in grade st	5 who are enrolled exactly four years later in
a public school in Grade 9)	

		Coho	rt	
	1997	1998	1999	2000
Female	0.5167	0.5127	0.5111	0.5132
White	0.5408	0.5481	0.5334	0.5249
Black	0.2327	0.2354	0.2344	0.2267
Hispanic	0.1985	0.1856	0.1999	0.2138
Asian	0.0211	0.023	0.0229	0.0231
Race-other	0.0069	0.008	0.0095	0.0114
Free Lunch	0.3985	0.3883	0.3896	0.384
Reduced-Price Lunch	0.0927	0.1035	0.1048	0.1091
5th-grade SSS Normed Math Score	0.0881	0.0797	0.0851	0.0843
5th-grade NRT Normed Math Score			0.0857	0.0925
6th-grade SSS Normed Math Score	-0.2911	-0.5562	0.1345	0.1462
6th-grade NRT Normed Math Score	-0.6128	0.1426	0.133	0.1457
7th-grade SSS Normed Math Score	-0.7875	0.1486	0.1361	0.1449
7th-grade NRT Normed Math Score	0.1309	0.1401	0.1354	0.145
8th-grade SSS Normed Math Score	0.1256	0.1375	0.1296	0.1371
8th-grade NRT Normed Math Score	0.1165	0.1309	0.1185	0.1277
9th-grade SSS Normed Math Score	0.1933	0.1797	0.1768	0.1682
9th-grade NRT Normed Math Score	0.1798	0.1628	0.1581	0.1501
10th-grade SSS Normed Math Score	0.1232	0.2457	0.2143	0.2327
10th-grade NRT Normed Math Score	0.0916	0.0956	0.0884	0.1077
Earned Regular HS Diploma within 4 years of Entering Grade 9	0.6799	0.6537	0.6652	0.6833
Earned Regular HS Diploma within 5 years of Entering Grade 9	0.6823	0.6563	0.6686	0.6867

Race/ Ethnicity	Math Test Scores							
	SSS Grade 5	NRT Grade 5	SSS Grade 6	NRT Grade 6	SSS Grade 7	NRT Grade 7	SSS Grade 8	NRT Grade 8
White	0.3190	0.3308	0.3674	0.4028	0.3633	0.3868	0.3413	0.3731
Black	-0.4070	-0.3815	-0.3276	-0.3609	-0.3344	-0.3611	-0.3532	-0.3879
Hispanic	-0.0230	-0.0502	0.0400	-0.0186	0.0440	0.0048	0.0762	-0.0041
Asian	0.5680	0.6522	0.7132	0.7832	0.7268	0.8221	0.7659	0.7749
Other	0.0680	0.0895	0.1730	0.1600	0.1524	0.1408	0.1525	0.1307

Table 3A - Normed Test Scores in Grades 5-8 by Race (Students with a test score in grade 5 in 2000 who are enrolled four years later in a public school in Grade 9)

Table 3B - Normed Test Scores in Grades 5-8 by Gender<br/>(Students with a test score in grade 5 in 2000 who are enrolled four years later<br/>in a public school in Grade 9)

Gender		Math Test Scores							
	SSS	NRT	SSS	NRT	SSS	NRT	SSS	NRT	
	Grade 5	Grade 5	Grade 6	Grade 6	Grade 7	Grade 7	Grade 8	Grade 8	
Male	0.1060	0.0931	0.1425	0.1641	0.1482	0.1630	0.1183	0.1345	
Female	0.0639	0.0919	0.1496	0.1284	0.1418	0.1283	0.1546	0.1214	

Table 3C - Normed Test Scores in Grades 5-8 by 5th-grade Lunch Status (Students with a test score in grade 5 in 2000 who are enrolled four years later in a public school in Grade 9)

Lunch Category	Math Test Scores							
	SSS	NRT	SSS	NRT	SSS	NRT	SSS	NRT
	Grade 5	Grade 5	Grade 6	Grade 6	Grade 7	Grade 7	Grade 8	Grade 8
Free	-0.3164	-0.3166	-0.2620	-0.2899	-0.2614	-0.2923	-0.2653	-0.3096
<b>Reduced Price</b>	0.0218	0.0041	0.0649	0.0457	0.0602	0.0348	0.0573	0.0090
Full Price	0.3977	0.4178	0.4669	0.4912	0.4638	0.4906	0.4490	0.4730

### Table 4A – High School Exit by Race (Students with a test score in grade 5 in 1997-2000 who are enrolled four years later in a public school in Grade 9)

Exit Reasons	White	Black	Hispanic	Asian	Other	Total
No Exit (Enrolled in	188,875	73,144	65,797	9,585	2,725	340,126
Each of Grades 9-12)	[69.67]	[62.31]	[65.14]	[83.76]	[59.05]	[67.27]
Dropped out	15,142	12,571	9,874	312	388	38,287
	[5.59]	[10.71]	[9.78]	[2.73]	[8.41]	[7.57]
Exit to home school	4,619	350	650	51	78	5,748
	[1.70]	[0.30]	[0.64]	[0.45]	[1.69]	[1.14]
Exit to private school	5,429	2,174	2,649	107	110	10,469
	[2.00]	[1.85]	[2.62]	[0.93]	[2.38]	[2.07]
Exit – other	57,050	29,153	22,041	1,389	1,314	110,947
	[21.04]	[24.83]	[21.82]	[12.14]	[28.47]	[21.94]

### Table 4B – High School Exit by Gender

(Students with a test score in grade 5 in 1997-2000 who are enrolled four years later in a public school in Grade 9)

Exit Reasons	Male	Female	Total
No Exit (Enrolled in	162,205	184,154	346,359
Each of Grades 9-12)	[64.87]	[69.87]	[67.44]
Dropped out	20,590	18,038	38,628
	[8.23]	[6.84]	[7.52]
Exit to home school	2,488	3,323	5,811
	[1.00]	[1.26]	[1.13]
Exit to private school	5,380	5,226	10,606
	[2.15]	[1.98]	[2.06]
Exit – other	59,381	52,832	112,213
	[23.75]	[20.04]	[21.85]

### Table 4C – High School Exit by Lunch Status

(Students with a test score in grade 5 in 1997-2000 who are enrolled four years later in a public school in Grade 9)

Exit Reasons	Free Lunch	Reduced Price	Full Price	Total
No Exit (Enrolled in	110,843	34,893	200,624	346,360
Each of Grades 9-12)	[55.64]	[65.77]	[76.76]	[67.44]
Dropped out	24,848	3,907	9,873	38,628
	[12.47]	[7.36]	[3.78]	[7.52]
Exit to home school	2,126	710	2,975	5,811
	[1.07]	[1.34]	[1.14]	[1.13]
Exit to private school	4,312	987	5,307	10,606
	[2.16]	[1.86]	[2.03]	[2.06]
Exit – other	57,077	12,559	42,577	112,213
	[28.65]	[23.67]	[16.29]	[21.85]

# Table 5A – Regular High School Diploma Receipt within 5 Years of Entering Grade 9 by Race

(Students with a test score in grade 5 in 1997-2000 who are enrolled four years later in a public school in Grade 9)

Diploma Received	White	Black	Hispanic	Asian	Other	Total
No	80,914	49,033	32,926	1,682	1,921	166,476
	[29.84]	[41.77]	[32.6]	[14.70]	[41.63]	[32.93]
Yes	190,201	68,359	68,085	9,762	2,694	339,101
	[70.16]	[58.23]	[67.4]	[85.30]	[58.37]	[67.07]

Note: numbers in brackets are column percentages.

# Table 5B - Regular High School Diploma Receipt within 5 Years of Entering Grade 9 by Gender

(Students with a test score in grade 5 in 1997-2000 who are enrolled four years later in a public school in Grade 9)

Diploma Received	Male	Female	Total
No	89,301	78,625	167,926
	[35.71]	[29.83]	[32.69]
Yes	160,743	184,948	345,691
	[64.29]	[70.17]	[67.31]

Note: numbers in brackets are column percentages.

# Table 5C – Regular High School Diploma Receipt within 5 Years of Entering Grade 9 by Lunch Status in Grade 5

(Students with a test score in grade 5 in 1997-2000 who are enrolled four years later in a public school in Grade 9)

Diploma Received	Free Lunch	<b>Reduced Price</b>	Full Price	Total
No	93,478	18,170	56,278	167,926
NO	[46.93]	[34.25]	[21.53]	[32.69]
Yes	105,728	34,886	205,078	345,692
105	[53.07]	[65.75]	[78.47]	[67.31]

#### Table 6A – Normed Test Scores in Grades 5-10 by Race

(Students with a test score in grade 5 in 2000 who are enrolled four years later in a public school in Grade 9 and continue to be enrolled in Grades 10 and 11)

Race		Math Test Scores										
	SSS Grade 5	NRT Grade 5	SSS Grade 6	NRT Grade 6	SSS Grade 7	NRT Grade 7	SSS Grade 8	NRT Grade 8	SSS Grade 9	NRT Grade 9	SSS Grade 10	NRT Grade 10
White	0.3692	0.3889	0.4241	0.4636	0.4217	0.4515	0.4029	0.4425	0.4325	0.4237	0.3857	0.3194
Black	-0.3452	-0.3251	-0.2630	-0.3057	-0.2694	-0.3037	-0.2820	-0.3295	-0.2407	-0.2781	-0.1249	-0.3406
Hispanic	0.0454	0.016	0.1099	0.0475	0.1161	0.0729	0.1524	0.0669	0.1731	0.1085	0.2253	0.0248
Asian	0.5890	0.6791	0.7345	0.8120	0.7518	0.8525	0.7911	0.8026	0.8364	0.9525	0.7264	0.8733
Other	0.1201	0.1532	0.2481	0.2329	0.2232	0.2148	0.2336	0.205	0.252	0.2421	0.2572	0.1321

#### Table 6B – Normed Test Scores in Grades 5-10 by Gender

(Students with a test score in grade 5 in 2000 who are enrolled four years later in a public school in Grade 9 and continue to be enrolled in Grades 10 and 11)

Gender		Math Test Scores										
	SSS Grade 5	NRT Grade 5	SSS Grade 6	NRT Grade 6	SSS Grade 7	NRT Grade 7	SSS Grade 8	NRT Grade 8	SSS Grade 9	NRT Grade 9	SSS Grade 10	NRT Grade 10
Male	0.1761	0.1673	0.2179	0.2397	0.2249	0.2407	0.1993	0.2167	0.2738	0.2379	0.2856	0.1444
Female	0.1138	0.1427	0.2027	0.1811	0.1966	0.1849	0.2126	0.1798	0.2032	0.1891	0.2409	0.1030

Lunch Category	Math Test Scores											
	SSS Grade 5	NRT Grade 5	SSS Grade 6	NRT Grade 6	SSS Grade 7	NRT Grade 7	SSS Grade 8	NRT Grade 8	SSS Grade 9	NRT Grade 9	SSS Grade 10	NRT Grade 10
Free	-0.2519	-0.2542	-0.1918	-0.2278	-0.1882	-0.2256	-0.1832	-0.2377	-0.1465	-0.2034	-0.0586	-0.2905
Reduced Price	0.062	0.0489	0.1162	0.0945	0.1103	0.0846	0.1158	0.0642	0.1421	0.0821	0.1679	-0.0244
Full Price	0.4264	0.451	0.4986	0.526	0.4968	0.5287	0.4833	0.5135	0.5096	0.5087	0.4578	0.4120

 Table 6C – Normed Test Scores in Grades 5-10 by Lunch Status in Grade 5

(Students with a test score in grade 5 in 2000 who are enrolled four years later in a public school in Grade 9 and continue to be enrolled in Grades 10 and 11)

Table 7A - College Attendance in Year Immediately Following Receipt of Regular High<br/>School Diploma by Race<br/>(Students with a test score in grade 5 in 1997 who are enrolled four years later<br/>in a public school in Grade 9 and graduate within 5 years with a regular high<br/>school diploma)

			Race		
	White	Black	Hispanic	Asian	Total
No College	11,165	4,828	3,741	274	20,164
	[31.06]	[37.86]	[29.77]	[15.77]	[31.82]
FL Community College	12,686	3,343	2,496	441	19,081
	[35.29]	[26.22]	[19.86]	[25.37]	[30.11]
4-year FL Public University	9,054	2,540	3,147	797	15605
	[25.19]	[19.92]	[25.04]	[45.86]	[24.62]
4-year FL Private College/Univ.	1,065	449	423	84	2,029
	[2.96]	[3.52]	[3.37]	[4.83]	[3.2]
4-year College Out of State	1,976	1,592	2,759	142	6,499
	[5.5]	[12.48]	[21.96]	[8.17]	[10.25]
Total	35,946	12,752	12,566	1,738	63,378
	[100]	[100]	[100]	[100]	[100]

Table 7B - College Attendance in Year Immediately Following Receipt of Regular High<br/>School Diploma by Gender<br/>(Students with a test score in grade 5 in 1997 who are enrolled four years later<br/>in a public school in Grade 9 and graduate within 5 years with a regular high<br/>school diploma)

	Ger	nder	
	Male	Female	Total
No College	10,610	9,802	20,412
	[35.72]	[28.02]	[31.56]
FL Community College	8,431	11,176	19,607
	[28.38]	[31.95]	[30.31]
4-year FL Public University	6,519	9,402	15,921
	[21.95]	[26.88]	[24.61]
4-year FL Private College/Univ.	934	1,139	2,073
	[3.14]	[3.26]	[3.2]
4-year College Out of State	3,210	3,460	6,670
	[10.81]	[9.89]	[10.31]
Total	29,704	34,979	64,683
	[100.00]	[100.00]	[100.00]

Table 7C – College Attendance in Year Immediately Following Receipt of Regular High<br/>School Diploma by Lunch Status in Grade 5<br/>(Students with a test score in grade 5 in 1997 who are enrolled four years later<br/>in a public school in Grade 9 and graduate within 5 years with a regular high<br/>school diploma)

		Lunch	Status	
	Free Lunch	Reduced Price	Full Price	Total
No College	8,983	2,215	9,214	20,412
	[44.58]	[37.75]	[23.83]	[31.56]
FL Community College	4,952	1,829	12,826	19,607
	[24.58]	[31.17]	[33.17]	[30.31]
4-year FL Public University	2,899	1,079	11,943	15,921
	[14.39]	[18.39]	[30.89]	[24.61]
4-year FL Private College/Univ.	573	155	1,345	2,073
	[2.84]	[2.64]	[3.48]	[3.2]
4-year College Out of State	2,743	590	3,337	6,670
	[13.61]	[10.05]	[8.63]	[10.31]
Total	20,150	5,868	38,665	64,683
	[100.00]	[100.00]	[100.00]	[100.00]

 Table 8 – Probit Estimates of the Determinants of Taking One or More Courses in a STEM

 Field in the First Year in College

 (Students with a test score in grade 5 in 1997-2000 who are enrolled four years

 later in a public school in Grade 9 and graduate within 5 years with a regular

 high school diploma and attend a Florida public university within one year of

 receiving their diploma)

Explanatory	(1)	(2)	(3)	(4)
Variables	Taking One or	Taking One or	Taking One or	Taking One or
	More STEM	More STEM	More STEM	More STEM
	Courses	Courses	Courses	Courses
Black	0.0320**	0.0338**	0.0323**	0.0324**
Diack	(0.0036)	(0.0039)	(0.0042)	(0.0046)
Hispanic	0.0540**	0.0553**	0. 0557**	0.0570**
Hispanic	(0.0033)	(0.0034)	(0.0035)	(0.0037)
Asian	0.0293**	0. 0301**	0.0283**	0.0267**
Asian	(0.0056)	(0.0057)	(0.0058)	(0.0060)
Other Race	0.0172	0.0177	0.0096	0.0109
Other Race	(0.0177)	(0.0177)	(0.0189)	(0.0195)
Es ma la	-0.0370**	-0.0369**	-0.0368**	-0.0360**
Female	(0.0027)	(0.0028)	(0.0029)	(0.0030)
Free Lunch		-0.0031	-0.0028	-0.0031
Flee Lunch		(0.0043)	(0.0044)	(0.0046)
Deduced Lynch		0088	-0.0091	-0.0110
Reduced Lunch		(0.0058)	(0.0059)	(0.0061)
		()	0.0027	0.0027
SSS Grade 9			(0.0040)	(0.0043)
NDT Crede 0			-0.0009	0.0007
NRT Grade 9			(0.0029)	(0.0031)
SSS Crede 10			-0.0111*	-0.0097
SSS Grade 10			(0.0051)	(0.0055)
			0.0051	0.0049
NRT Grade 10			(0.0027)	(0.0028)
				-0.00136
NRT Grade 7				(0.0030)
				0.0060
SSS Grade 8				(0.0050)
				-0.0042
NRT Grade 8				(0.0033)
Observations	74,528	74,528	71,113	67,424
Psuedo R <sup>2</sup>	0.0090	0.0090	0.0091	0.0090

Note: All models include cohort controls. Reported estimates are marginal effects. Standard errors in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 9 – Probit Estimates of the Determinants of Taking One or More Courses in SpecificScience Subjects During the First Year in College (Students with a test score ingrade 5 in 1997-2000 who are enrolled four years later in a public school in Grade9 and graduate within 5 years with a regular high school diploma and attend aFlorida public university within one year of receiving their diploma)

Explanatory Variables	Taking One or More Science Courses	Taking One or More Biology Courses	Taking One or More Chemistry Courses	Taking One or More Physics Courses
Black	0.0527**	0.0208**	0.1212**	-0.0032
Баск	(0.0061)	(0.0055)	(0.0061)	(0.0020)
Hisponia	0.0982**	0.0304**	0.1177**	0.0054**
Hispanic	(0.0052)	(0.0049)	(0.0052)	(0.0017)
A	0.0721**	0.0749**	0.1626**	0.0097**
Asian	(0.0082)	(0.0081)	(0.0084)	(0.0027)
Other Dees	0.0000	-0.0231	0.0398	-0.0024
Other Race	(0.0267)	(0.0229)	(0.0251)	(0.0074)
Esmals	-0.0110**	0.0616**	-0.0048	-0.0312**
Female	(0.0040)	(0.0034)	(0.0034)	(0.0014)
Free Lunch	-0.0117*	-0.0099*	0.0085	0.0042*
	(0.0059)	(0.0050)	(0.0050)	(0.0020)
D 1 17 1	-0.0165*	-0.0085	0.0024	0.0000
Reduced Lunch	(0.0078)	(0.0066)	(0.0066)	(0.0024)
	-0.0029	-0.0174**	0.0248**	0.0055**
SSS Grade 9	(0.0056)	(0.0051)	(0.0047)	(0.0013)
	0.0048	0.0030	0.0055	0.0029**
NRT Grade 9	(0.0041)	(0.0036)	(0.0034)	(0.0010)
000 0 1 10	0.0170*	-0.0157*	0.0631**	0.0158**
SSS Grade 10	(0.0073)	(0.0065)	(0.0061)	(0.0017)
	0.0083*	-0.0046	0.0227**	0.0100**
NRT Grade 10	(0.0037)	(0.0033)	(0.0031)	(0.0009)
	-0.0100*	-0.0097**	-0.0068*	0.0008
NRT Grade 7	(0.0040)	(0.0035)	(0.0034)	(0.0011)
	0.0128*	-0.0046	0.0148**	0.0047**
SSS Grade 8	(0.0065)	(0.0058)	(0.0055)	(0.0016)
	-0.0137**	-0.0087*	-0.0136**	0.0004
NRT Grade 8	(0.0043)	(0.0038)	(0.0037)	(0.0011)
Observations	67,424	67,424	67,424	67,424
Psuedo R <sup>2</sup>	0.0060	0.0144	0.0294	0.1401

Note: All models include cohort controls. Reported estimates are marginal effects. Standard errors in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 10 -Probit Estimates of the Determinants of Taking One or More Courses in a<br/>STEM Field During the First Year in College<br/>(Students with a test score in grade 5 in 1997-2000 who are enrolled four years<br/>later in a public school in Grade 9 and graduate within 5 years with a regular<br/>high school diploma and attend a Florida public university within one year of<br/>receiving their diploma)

Explanatory Variables	(1) Taking One or	(2) Taking One or	(3) Taking One or
	More STEM	More STEM	<b>More STEM</b>
	Courses	Courses	Courses
Black	0.0470**	0.0302**	0.0450**
Black	(0.0053)	(0.0049)	(0.0061)
Hispanic	0.0324**	0.0540**	0.0280**
Hispanic	(0.0048)	(0.0042)	(0.0056)
Asian	0.0277**	0.0202**	0.0209**
Asian	(0.0058)	(0.0070)	(0.0069)
Other Race	0.0166	0.0088	0.0087
Other Race	(0.0185)	(0.0215)	(0.0214)
Female	-0.0745**	-0.0408**	-0.0700**
	(0.0071)	(0.0034)	(0.0087)
Female x Proportion of Middle			
and High School Math and	0.0653**		0.0514**
Science Courses Taught by a	(0.0116)		(0.0142)
Female Teacher			
Black x Proportion of Middle			
and High School Math and	-0.0648**		-0.0692**
Science Courses Taught by a	(0.0206)		(0.0234)
Black Teacher			
Hispanic x Proportion of			
Middle and High School Math	0.1560**		0.1640**
and Science Courses Taught	(0.0185)		(0.0123)
by a Hispanic Teacher			
Enrolled in at Least One HS		0.0132*	0.0124*
Biology Course Taught by a		(0.0059)	(0.0059)
Teacher with a BA in Biology		(0.0057)	(0.0000))
Enrolled in at Least One HS			
Chemistry Course Taught by a		0.0236**	0.0214**
Teacher with a BA in		(0.0073)	(0.0074)
Chemistry			
Enrolled in at Least One HS		-0.0104	-0.0118
Physics Course Taught by a		(0.0121)	(0.0122)
Teacher with a BA in Physics		(0.01-1)	(0.01)
Enrolled in at Least One HS		0.0134**	0.0144**
Math Course Taught by a		(0.0051)	(0.0050)
Teacher with a BA in Math			. ,
Observations	69,978	49,639	49,634
Psuedo R <sup>2</sup>	0.0108	0.0103	0.0121

Note: All models include cohort controls, controls for free and reduced-price lunch in grade 5 and controls for math test scores in grades 5, 7 and 8. Reported estimates are marginal effects. Standard errors are in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 11 – Probit Estimates of the Determinants of Taking One or More Courses in<br/>Specific Science Subjects During the First Year in College<br/>(Students with a test score in grade 5 in 1997-2000 who are enrolled four years<br/>later in a public school in Grade 9 and graduate within 5 years with a regular<br/>high school diploma and attend a Florida public university within one year of<br/>receiving their diploma)

Explanatory Variables	Taking One or More Science Courses	Taking One or More Biology Courses	Taking One or More Chemistry Courses	Taking One or More Physics Courses	
	-0.0457**	0.0252**	0.1082**	-0.0054	
Black	(0.0069)	(0.0064)	(0.0069)	(0.0022)	
<b>TT</b>	-0.0854**	0.0219**	0.1147**	0.0046	
Hispanic	(0.0060)	(0.0057)	(0.0060)	(0.0020)	
<b>A</b> = <b>i</b> =	0.0732**	0.0708**	0.1738**	0.0139	
Asian	(0.0095)	(0.0095)	(0.0099)	(0.0034)	
Other Race	-0.0134	-0.0120	0.0198	-0.0004	
Other Race	(.0298)	(0.0263)	(0.0276)	(0.0087)	
Famala	-0.0141**	0.0678**	-0.0140**	-0.0342	
Female	(0.0046)	(0.0040)	(0.0040)	(0.0017)	
Enrolled in at Least One HS	0.0082	-0.0069	0.0220**	0.0028	
Biology Course Taught by a	(0.0081)	(0.0071)	(0.0072)	(0.0025)	
Teacher with a BA in Biology	(0.0001)	(0.0071)	(0.0072)	(0.0023)	
Enrolled in at Least One HS					
Chemistry Course Taught by a	-0.0216*	0.0008	0.0185*	0.0022	
Teacher with a BA in	(0.0101)	(0.0090)	(0.0090)	(0.0032)	
Chemistry					
Enrolled in at Least One HS	-0.0027	-0.0070	0.0251*	0.0032	
Physics Course Taught by a	(0.0155)	(0.0137)	(0.0136)	(0.0044)	
Teacher with a BA in Physics	(0.0155)	(0.0157)	(0.0150)	(0.0011)	
Enrolled in at Least One HS	0.0393**	0.0159**	0.0210**	0.0083	
Math Course Taught by a	(0.0068)	(0.0062)	(0.0061)	(0.0023)	
Teacher with a BA in Math	(0.000)	(0.0002)	(0.0001)	(0.0025)	
Observations	49,639	49,639	49,639	49,639	
Psuedo R <sup>2</sup>	0.0053	0.0138	0.0256	0.1280	

Note: All models include cohort controls, controls for free and reduced-price lunch in grade 5 and controls for math test scores in grades 5, 7 and 8. Reported estimates are marginal effects. Standard errors are in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 12-Probit Estimates of the Determinants of Earning a Bachelor's Degree in a<br/>STEM Major Within 9 years of Starting Grade 9 - No Pre-College Controls<br/>(Students with a test score in grade 5 in 1997 who are enrolled four years later<br/>in a public school in Grade 9 and graduate within 5 years with a regular high<br/>school diploma and attend a Florida public university within one year of<br/>receiving their diploma)

Explanatory Variables	Earning a Bachelor's Degree in STEM	Earning a Bachelor's Degree in Math	Earning a Bachelor's Degree in Chemistry	Earning a Bachelor's Degree in Physics	Earning a Bachelor's Degree in Bio. Sci.	Earning a Bachelor's Degree in Engineering
Black	-0. 0377**	-0.0019**	-0.0021**		-0.0031**	-0.0180**
DIRCK	(0.0020)	(0.0004)	(0.0006)		(0.0005)	(0.0010)
Hispanic	-0.0097**	-0.0016**	-0.0003	-0.0004	-0.0028**	-0.0006**
Inspanc	(0.0022)	(0.0004)	(0.0006)	(0.0002)	(0.0005)	(0.0011)
Asian	0.0491**	0.0007	0.0077**	0.0002	-0.0005	0.0067**
Asian	(0.0049)	(0.0008)	(0.0017)	(0.0005)	(0.0010)	(0.0023)
Other Race	-0.0110	0.0008	0.0022		0.0018	-0.0056
Ould Race	(0.0107)	(0.0026)	(0.0040)		(0.0036)	(0.0054)
Female	-0.0497**	-0.0021**	-0.0022**	-0.0030**	0.0016**	-0.0416**
remate	(0.0020)	(0.0004)	(0.0005)	(0.0004)	(0.0005)	(0.0014)
Observations	75,292	75,292	75,292	62,816	75,292	75,292
Psuedo R <sup>2</sup>	0.0335	0.0198	0.0178	0.0603	0.0163	0.0782

Note: All models include cohort controls. Reported estimates are marginal effects. Shaded cells represent variables that were dropped from the estimated model due to the fact that they perfectly predicted success or failure. Standard errors are in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 13 – Probit Estimates of the Determinants of Earning a Bachelor's Degree in a<br/>STEM Major Within 9 years of Starting Grade 9 – With Pre-College Controls<br/>(Students with a test score in grade 5 in 1997 who are enrolled four years later<br/>in a public school in Grade 9 and graduate within 5 years with a regular high<br/>school diploma and attend a Florida public university within one year of<br/>receiving their diploma)

Explanatory Variables	Earning a Bachelor's Degree in STEM	Earning a Bachelor's Degree in Math	Earning a Bachelor's Degree in Chemistry	Earning a Bachelor's Degree in Physics	Earning a Bachelor's Degree in Bio. Sci.	Earning a Bachelor's Degree in Engineering
Black	0.00079	0.0006	0.0009		-0.0014	-0.0012
Didek	(0.0031)	(0.0006)	(0.0009)	(0.0007 0.0002 -0.002	(0.0007)	(0.0015)
Hispanic	0.0072**	-0.0003	0.0009	0.0002	-0.0023	0.0015
mspanic	(0.0025)	(0.0003)	(0.0007)	(0.0002)	2 -0.0023 ) (0.0006) -0.0006	(0.0011)
Asian	0.0405**	0.0002	0.0042**	0.0001	-0.0006	0.0027
Asian	(0.0046)	(0.0005)	(0.0013)	(0.0003)	(0.0009)	(0.0016)
Other Race	0.0016	0.0014	0.0026		-0.0009	0.0020
Other Race	(0.0121)	(0.0024)	(0.0004)		-0.0023 (0.0006) -0.0006 (0.0009) -0.0009 (0.0029) 0.0022** (0.0005)	(0.0057)
Female	-0.0262	-0.0003	-0.0004	-0.0011**	0.0022**	-0.0028**
remale	(0.0018)	(0.0002)	(0.0004)	(0.0002)	(0.0005)	(0.0011)
Observations	68,093	68,093	68,093	56,883	68,093	68,093
Psuedo R <sup>2</sup>	0.1093	0.1319	0.0760	0.1464	0.0284	0.2631

Note: All models include cohort controls, controls for free and reduced-price lunch in grade 5 and controls for math test scores in grades 7, 8, 9 and 10. Reported estimates are marginal effects. Shaded cells represent variables that were dropped from the estimated model due to the fact that they perfectly predicted success or failure. Standard errors are in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 14 – Probit Estimates of the Determinants of Earning a Bachelor's Degree in a<br/>STEM Major Within 9 years of Starting Grade 9<br/>(Students with a test score in grade 5 in 1997-2000 who are enrolled four years<br/>later in a public school in Grade 9 and graduate within 5 years with a regular<br/>high school diploma and attend a Florida public university within one year of<br/>receiving their diploma)

Explanatory Variables	Earning a Bachelor's	Earning a Bachelor's	Earning a Bachelor's	Earning a Bachelor's	Earning a Bachelor's	Earning a Bachelor's
	Degree in	Degree in	Degree in	Degree in	Degree in	Degree in
	STEM	Math	Chemistry	Physics	Bio. Sci.	<b>Engineering</b>
Black	-0.0060*	0.0004	0.0001		-0.0015*	-0.0028*
DIACK	(0.0027)	(0.0005)	(0.0006)		Bachelor's Degree in Bio. Sci. -0.0015* (0.0006) -0.0024** (0.0005) -0.0013 (0.0007) -0.0010 (0.0024) 0.0017** (0.0005) -0.0019 (0.0017) -0.0019** (0.0010) -0.0004 (0.0003) -0.0014* (0.0006) 0.0013** (0.0001) 67,424	(0.0011)
Hispanic	-0.0022	-0.0003	0.0002	0.0000	-0.0024**	-0.0004
Inspanie	(0.0022)	(0.0003)	(0.0005)	(0.0001)       (0.0005)         *       -0.0000       -0.0013         (0.0002)       (0.0007)         -0.0010       (0.0024)         -0.0007**       0.0017**         (0.0002)       (0.0005)         0.0000       -0.0019         (0.0001)       (0.0017)	(0.0009)	
Asian	0.0277**	0.0002	0.0021**	-0.0000	-0.0013	0.0020
7 (Shtill	(0.0041)	(0.0004)	(0.0009)	$\begin{array}{c c} & -0.0015^{*} \\ (0.0006) \\ 0.0000 \\ -0.0024^{**} \\ (0.0001) \\ (0.0005) \\ -0.0000 \\ -0.0013 \\ (0.0002) \\ (0.0007) \\ -0.0010 \\ (0.00024) \\ -0.0010 \\ (0.00024) \\ -0.0010 \\ (0.00024) \\ -0.0010 \\ (0.0005) \\ 0.0000 \\ -0.0019 \\ (0.0001) \\ (0.0017) \\ -0.0002^{*} \\ (0.0019 \\ (0.0019) \\ (0.0019 \\ (0.0001) \\ (0.0001) \\ (0.0001) \\ (0.0003) \\ \end{array}$	(0.0007)	(0.0013)
Other Race	0.0006	0.0012	0.0016			0.0010
Other Race	(0.0111)	(0.0022)	(0.0031)			(0.0045)
Female	-0.0141**	-0.0002	-0.0005	-0.0007**	0.0017**	-0.0125**
	(0.0017)	(0.0002)	(0.0003)	(0.0002)	(0.0005)	(0.0009)
No. of Computer	0.0250**	0.0004*	-0.0023	0.0000	0.0010	0.0002
Courses in First	(0.0022)	(0.0002)	(0.0012)		0.00022	(0.0002)
Year	(0.0022)	(0.0002)	(0.0012)	(0.0001)	(0.0017)	(0.0008)
No. of Engineer-						
ing Courses in	0.0201**	-0.0005**	-0.0014**			0.0087**
First Year	(0.0010)	(0.0002)	(0.0004)	(0.0001)	(0.0010)	(0.0004)
No. of Math						
Courses in First	0.0058**	0.0008**	0.0003	0.0001**	-0.0004	0.0027**
Year	(0.0009)	(0.0001)	(0.0002)	(0.0001)	(0.0003)	(0.0004)
No. of Statistics						
	-0.0312**	0.0003	-0.0019**	-0.0006**	-0.0014*	-0.0140**
Courses in First	(0.0004)	(0.0002)	(0.0005)	(0.0002)	(0.0006)	(0.0012)
Year	· · · · ·			~ /	× /	· · · ·
No. of Science	0.0125**	-0.0000	0.0010**	0.0001**	0.0012**	0.0012**
Courses in First	(0.0004)	(0.0001)	(0.0001)	(0.0000)		(0.0002)
Year	(0.0004)	(0.0001)	(0.0001)	(0.000)	(0.0001)	(0.0002)
Observations	67,424	67,424	67,424	56,401	67,424	50,422
Psuedo R <sup>2</sup>	0.1665	0.1511	0.1295	0.2010	0.0599	0.2735

Note: All models include cohort controls, controls for free and reduced-price lunch in grade 5 and controls for math test scores in grades 7, 8, 9 and 10. Reported estimates are marginal effects. Shaded cells represent variables that were dropped from the estimated model due to the fact that they perfectly predicted success or failure. Standard errors are in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.

Table 15 – Probit Estimates of the Determinants of Earning a Bachelor's Degree in a<br/>STEM Major Within 9 years of Starting Grade 9<br/>Students with a test score in grade 5 in 1997-2000 who are enrolled four years<br/>later in a public school in Grade 9 and graduate within 5 years with a regular<br/>high school diploma and attend a Florida public university within one year of<br/>receiving their diploma)

Explanatory Variables	Earning a Bachelor's Degree in					
	STEM	Math	Chemistry	Physics	Bio. Sci.	Engineering
Black	0.0004	0.0005	0.0013		-0.0003	-0.0028
	(0.0041)	(0.0006)	(0.0010)	0.0001	(0.0010)	(0.0015)
Hispanic	-0.0003	-0.0007*	-0.0006	-0.0001	-0.0012	0.0007
1	(0.0028)	(0.0003)	(0.0005)	(0.0001)	(0.0007)	(0.0011)
Asian	0. 0268**	0.0002	0.0021**	-0.0000	-0.0013	0.0018
	(0.0040)	(0.0004)	(0.0008)	(0.0001)	(0.0006)	(0.0013)
Other Race	0.0000	0.0012	0.0017		-0.0008	0.0012
other Ruce	(0.0109)	(0.0021)	(0.0030)		(0.0023)	(0.0045)
Female	-0.0009	0.0004	0.0009	-0.0005	0.0020	-0.0095**
	(0.0047)	(0.0006)	(0.0008)	(0.0005)	(0.0011)	(0.0026)
Controls for						
First-Year	Yes	Yes	Yes	Yes	Yes	Yes
Coursework						
Female x Proportion of Middle and High School Math and Science Courses Taught by a Female Teacher	-0.0142* (0.0072)	-0.0007 (0.0003)	-0.0020 (0.0014)	-0.0004 (0.0006)	0.0006 (0.0017)	-0.0007 (0.0034)
Black x Proportion of Middle and High School Math and Science Courses Taught by a Black Teacher	-0.0332* (0.0161)	0.0000 (0.0021)	-0.0065 (0.0038)		-0.0070 (0.0043)	0.0045 (0.0072)
Hispanic x Proportion of Middle and High School Math and Science Courses Taught by a Hispanic Teacher Female x	0.0044 (0.0094)	0.0027* (0.0013)	0.0033* (0.0017)	-0.0000 (0.0006)	-0.0086* (0.0036)	-0.0005 (0.0038)
Proportion of First-Year College STEM Courses	-0.0190** (0.0039)	-0.0010 (0.0005)	-0.0009 (0.0008)	0.0002 (0.0003)	-0.0033** (0.0010)	-0.0068** (0.0019)

Taught by a						
Female Instructor						
Black x						
Proportion of						
First-Year College	-0.0662**	-0.0011	-0.0037		-0.0018	-0.0132
Courses Taught by	(0.0154)	(0.0020)	(0.0032)		(0.0033)	(0.0071)
a Black by a Black						
Instructor						
Hispanic x						
Proportion of						
First-Year College	-0.0210**	-0.0001	0.0001	0.0009*	-0.0002	-0.0068*
Courses Taught by	(0.0080)	(0.0012)	(0.0015)	(0.0004)	(0.0027)	(0.0031)
an Hispanic						
Instructor						
Observations	65,458	65,458	65,458	54,715	65,458	65,458
Psuedo R <sup>2</sup>	0.1727	0.1594	0.1388	0.2046	0.0692	0.2781

Note: All models include cohort controls, controls for free and reduced-price lunch in grade 5 and controls for math test scores in grades 7, 8, 9 and 10. Shaded cells represent variables that were dropped from the estimated model due to the fact that they perfectly predicted success or failure. Reported estimates are marginal effects. Standard errors are in parentheses. \* significant at the 5% level, \*\* significant at the 1% level in a two-tailed test.